















Redefining the risk of major arrhythmic events in non-ischaemic cardiomyopathy: insights from the DERIVATE-NICM study

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Aims

Selection of the patients for implantable cardioverter defibrillator primary prevention therapy in non-ischaemic cardiomyopathy (NICM) needs to be improved. To evaluate the additional prognostic value of a new cardiac magnetic resonance (CMR) score based on late gadolinium enhancement (LGE) pattern distribution (DERIVATE Risk Score 2.0) when compared with previously published DERIVATE Risk Score 1.0, which is based solely on quantitative parameters, in a cohort of NICM patients enrolled in the DERIVATE registry.

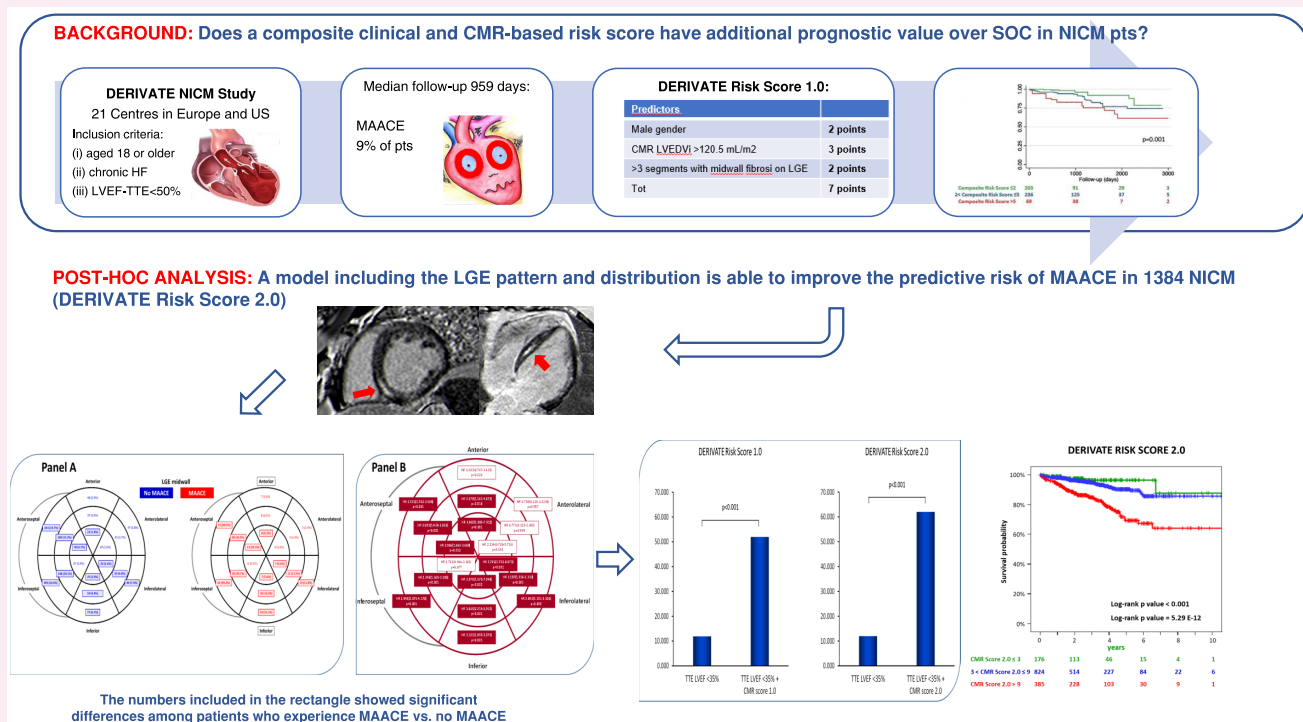
Methods and results

One thousand three hundred and eighty-four NICM patients with chronic heart failure and left ventricular ejection fraction (LVEF) < 50% were evaluated for primary sudden cardiac death prevention therapy. Major adverse arrhythmic cardiac events (MAACEs) were the primary endpoint. During a median follow-up of 959 days, MAACE occurred in 128 (9.2%) patients. In the multivariate analyses, male gender [hazard ratio (HR): 1.605 (95% confidence interval, CI: 1.051–2.451); $P = 0.028$], LVEF per point % [HR: 0.977 (95% CI: 0.961–0.993); $P = 0.005$] and presence and location of midwall LGE [weighted HR: 1.066 (95% CI: 1.045–1.086), $P < 0.001$] were independent predictors of MAACE. A multi-parametric CMR-weighted predictive-derived score (DERIVATE Risk Score 2.0) provided a higher additional prognostic value vs. transthoracic echocardiography–LVEF cut-off of 35% when compared with the previous published DERIVATE Risk Score 1.0 with a net reclassification improvement of 54.52% (95% CI: 36.52–72.52%; $P < 0.001$). These findings were confirmed in the validation cohort.

Conclusion

The presence of midwall LGE, but also the location of scar, confers an added and independent MAACE risk to a large NICM population influencing the choice of treatment.

Graphical Abstract



Keywords

non-ischaemic dilated cardiomyopathy • heart failure • cardiac magnetic resonance • implantable cardioverter defibrillator therapy

Introduction

Treatment with an implantable cardioverter defibrillator (ICD) has proven to be an effective prophylactic strategy for the prevention of sudden cardiac death (SCD) in patients with non-ischaemic dilated

cardiomyopathy (NICM).¹ During the past two decades, an increasing number of studies have shown that the international guidelines recommending ICD implantation lead to limited clinical benefit.² Indeed, the left ventricle (LV) ejection fraction (EF) centred strategy excludes many patients who would need a ICD protection and, at the same

Table 3 Univariable predictors of MAACE

	MAACE	
	HR (95% CI)	P-value
Demographic characteristics		
Age, years (per 1 year)	1.009 (0.996–1.022)	0.166
Male	1.714 (1.133–2.594)	0.011
Cardiovascular risk factor		
Family history	1.305 (0.909–1.872)	0.149
Smoking history	0.632 (0.420–0.952)	0.028
Hypertension	1.306 (0.922–1.849)	0.132
Hyperlipidaemia	0.960 (0.662–1.392)	0.830
Diabetes	1.287 (0.806–2.056)	0.290
NYHA Classes III and IV	1.228 (0.809–1.862)	0.334
Medical therapy		
Beta-blockers	3.161 (1.475–6.775)	0.003
Ivabradine	1.178 (0.515–2.693)	0.699
ACE inhibitors/AT1 blockers	0.981 (0.580–1.659)	0.943
Diuretics	1.524 (1.003–2.316)	0.049
Calcium-blockers	2.216 (1.218–4.033)	0.009
Anti-thrombotic agents	0.818 (0.551–1.214)	0.319
Anti-coagulant therapy	1.650 (1.085–2.510)	0.019
Nitrates	0.961 (0.447–2.066)	0.920
Statins	1.507 (1.041–2.182)	0.030
Amiodarone/other anti-arrhythmics	1.980 (1.308–2.996)	0.001
TTE		
LVEDVi (per 1 mL/m ²)	1.012 (1.007–1.016)	<0.001
LVESVi (per 1 mL/m ²)	1.014 (1.009–1.019)	<0.001
LVEF (per point %)	0.975 (0.959–0.992)	0.004
LVEF <35%	2.052 (1.354–3.110)	0.001
CMR functional evaluation		
LVEDVi (per 1 mL/m ²)	1.011 (1.008–1.014)	<0.001
LVEDVi >120.5 mL/m ²	3.238 (2.117–4.953)	<0.001
LVESVi (per 1 mL/m ²)	1.011 (1.008–1.015)	<0.001
LV mass (per 1 g/m ²)	1.003 (0.997–1.009)	0.295
LVSv (per 1 mL)	1.007 (1.001–1.013)	0.017
LVEF (per point %)	0.970 (0.954–0.985)	<0.001
LVEF <30%	2.168 (1.520–3.093)	<0.001
RVEDVi, mL/m ²	1.003 (0.998–1.008)	0.255
RVESVi, mL/m ²	1.004 (0.998–1.010)	0.178
RVSv, mL	1.003 (0.996–1.010)	0.401
RVEF, %	0.996 (0.982–1.009)	0.519
CMR LGE evaluation		
Prevalence of LGE-positive patients	2.650 (1.821–3.857)	<0.001
No. of segments with LGE	1.115 (1.073–1.159)	<0.001
Presence of midwall LGE pattern	2.644 (1.843–3.794)	<0.001
Presence of epicardial LGE pattern	1.319 (0.810–2.150)	0.266
Presence of mixed LGE pattern	1.670 (0.921–3.027)	0.091

For abbreviations, see Table 2.

Statistical method

Statistical analysis was performed using SPSS 25 (SPSS Inc.), R version 3.3, and Stata version 14 (StataCorp LLC). Continuous variables were expressed as mean \pm SD or median (25th–75th percentile) as appropriate and discrete variables as absolute numbers and percentages. Student's *t*-test or Mann–Whitney *U* test was used as appropriate to compare continuous variables between patients with and without MAACE while χ^2 test or Fisher's exact test was used to study differences regarding categorical data.

The overall analytic goal was to test whether a new DERIVATE Risk Score 2.0 including LGE location provides a higher additional prognostic value vs. current guidelines based on TTE–LVEF in order to predict MAACE when compared with the CMR-derived risk score model previously published and now referred to as DERIVATE Risk Score 1.0.⁶ Briefly, the DERIVATE Risk Score 1.0 comprised 7 points with 2 points assigned to male gender, 3 points to CMR LVEDVi > 120.5 mL/m², and 2 points to the presence of >3 segments with midwall fibrosis on LGE regardless the location of the scar.

In order to develop the DERIVATE Risk Score 2.0, we did the following steps, as previously described.¹² First, we evaluated the univariate Cox proportional hazard ratio (HR) of MAACE prediction of all baseline characteristics, including the presence of LGE for each myocardial segment. Second, we estimated the regression coefficients of the multiple logistic regression of variables significant at univariate Cox analysis after excluding collinear predictors. Third, we selected a referent risk factor profile by choosing a base category for each risk factor (the base category was the category assigned 0 points in the scoring system). Four, we computed how far each category of each risk factor is from the base category in terms of regression units. Five, we defined the number of regression units that will correspond to 1 point. Finally, we determined the risks that are associated with each point total.

The incremental value of CMR-DERIVATE Risk Score 1.0 and CMR-DERIVATE Risk Score 2.0 in addition to TTE–LVEF clinical data was assessed and compared by the χ^2 using Omnibus test of model coefficients.

Reclassification of patients of the new score was determined using net reclassification improvement (NRI) analysis for MAACE.

Finally, event-free survival related to the study endpoints was estimated using the Kaplan–Meier method, and survival curves were compared by means of the log-rank test among low, intermediate, and high-risk populations for MAACE according to the first, second/third, and fourth quartiles of CMR-DERIVATE Risk Score 2.0.

An independent validation cohort of patients who met the inclusion criteria of the study was subsequently identified to test the ability of the DERIVATE Risk Score 2.0 to adequately predict events in a different study population.

All results were considered significant with values of $P < 0.05$.

Results

According to the specified inclusion and exclusion criteria, the entire cohort consisted of 1384 subjects [mean age: 56 \pm 14 years, male: 948 (68.4%)]. Patient baseline characteristics are listed in Table 1. TTE and CMR tests were performed successfully in all patients with a median interval of 3 days (25th–75th percentile: 2–5 days) between TTE and CMR.⁶ The median follow-up time was 959 days (25th–75th percentile: 559.5–1590). MAACE occurred in 128 (9.2%).

Characteristics of the population according to the events

Patients who experienced MAACE were predominantly male, with no history of smoking and with more frequent use of beta-blockers, diuretics,

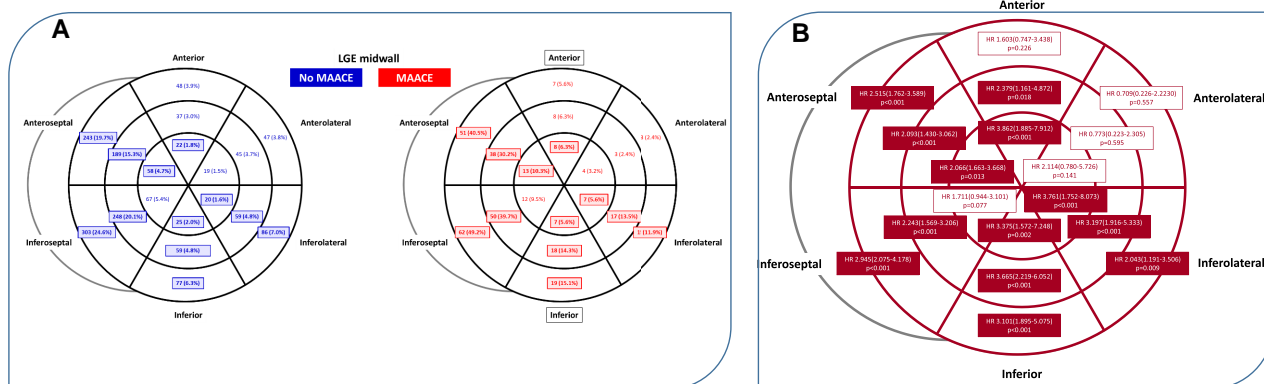


Figure 1 (A) Prevalence of midwall LGE in the LV myocardial segments between patient without MAACE (left) vs. patient who experienced MAACE (right). (B) HR to predict MAACE according to the LGE distribution. LGE, late gadolinium enhancement; MAACE, major arrhythmic adverse cardiac events. The numbers included in the rectangle showed significant differences among patients who experience MAACE vs. no MAACE.

Table 4 Multivariate predictors of MAACE to build DERIVATE Risk Score 2.0

	HR (95% CI)	P-value	Categories	Points
Male	1.605 (1.051–2.451)	0.028	Female	0
			Male	4
LVEF (per point %)	0.977 (0.961–0.993)	0.005	<24.9%	6
			25.0–34.9%	4
			35.0–44.9%	2
			>45.0%	0
LGE location	1.066 (1.045–1.086)	<0.001	0	0
—HR			>1 and ≤5	1
segment sum (per point)			>5 and ≤10	3
			>10	8

For abbreviations, see Table 2.

calcium-blockers, and anti-arrhythmics drugs when compared with the patients without MAACE (Table 1). Regarding imaging variable LVEDi, LVESi were higher and LVEF lower in MAACE group for both TTE and CMR when compared with patients without MAACE (Table 2). Moreover, the presence, the extent of LGE, and the midwall LGE pattern significantly differed among MAACE vs. non-MAACE (Table 2).

Predictors of MAACE

Regarding the imaging variables at univariate analysis (Table 3) LVEDi, LVESi and LVEF for both TTE and CMR were predictive of MAACE. Regarding the tissue characteristics, the presence, the extent of LGE, and the midwall LGE pattern also were predictive of MAACE. Regarding the role of midwall LGE distribution, Figure 1 shows the prevalence of LGE in each myocardial segment between patients who experienced MAACE vs. patients who had no events during the follow-up and the HR at univariate analyses for MAACE according to the LGE distribution.

According to the univariate analysis and after excluding collinear predictors, we built a new DERIVATE Risk Score 2.0 measuring the regression coefficients of the multiple logistic regression of variables significant at univariate Cox analysis, and we assigned a scoring for each variable (Table 4). Based on these results, the study population was divided into 4 quartiles by considering low (176 subjects), intermediate (824 subjects), and high-risk patients (385 subjects) who fell into the first (Q1), second to third (Q2 and Q3), and fourth quartiles (Q4), respectively (Figure 2). All results shown in bold are considered statistically significant.

The omnibus test of model coefficients showed a higher additional prognostic value vs. TTE–LVEF cut-off of 35% when compared with the previously published DERIVATE Risk Score 1.0 with an NRI of 54.52% [95% confidence interval (CI): 36.52–72.52%; $P < 0.001$] (Figure 3).

Additionally, as shown in Figure 4, 85 patients out of 1031 with less than three segments with middle-wall LGE who experienced MAACE, 35 and 58% had high and intermediate risk, respectively, based on the DERIVATE Risk Score 2.0. Conversely, among the 41 patients out of 201 with more than three segments with midwall LGE who experienced MAACE, the DERIVATE Risk Score 2.0 upgraded all patients classified as intermediate risk according to the DERIVATE Risk Score 1.0 to high risk.

The validation cohort consisted of 244 patients with baseline characteristics listed in Supplementary data online, Table S1. MAACE occurred in 19 (8%) cases. As for the derivation cohort, the DERIVATE Risk Score 2.0 provided an NRI for MAACE of 38.1% (95% CI: 20.82–55.37%; $P < 0.001$) when compared with TTE–LVEF with Kaplan–Meier showing again significantly different event-free rates ($P < 0.001$; Supplementary data online, Figure S1).

Figure 5 shows two case examples and the main findings of the study are synthesized in the Graphical Abstract.

Discussion

The present sub-analysis of the entire DERIVATE-NICM cohort population demonstrates that: (i) the prevalence of LGE in the anteroseptal, inferoseptal, inferior, and inferolateral wall is higher in patients who experienced MAACE, while no differences there were in terms of LGE presence in the anterior and anterolateral wall among patients with MAACE vs. patients without MAACE; (ii) a new multi-parametric score (DERIVATE-NICM Risk Score 2.0) weighted on multivariate analysis and including gender, LVEF, and LGE location provides incremental

continuous stretching in this region could be the trigger for susceptible tissue when LGE is present.

Finally, the contribution of LGE located in the anterior and anterolateral walls to MAACE is not significant. This could reflect a different aetiology of the condition, as anterior and anterolateral LGE might conceal cases of patients with undiagnosed myocarditis, which, as is known, has a better prognosis compared with true NICM.

Limitations

Firstly, the results of our registry are possibly affected by referring biases. However, the sites included in the registry represent referral centres, where the CMR is part of the usual care. Secondly, in this registry, a relatively low MAACE rate was observed, likely due to the inclusion of patients without a history of ventricular arrhythmias, and the exclusion of patients with hypertrophic cardiomyopathy and arrhythmogenic right ventricular cardiomyopathy, which was done by intention to investigate the prognostic yield of CMR in a lower risk population.

Furthermore, in the current analysis, we intended to ensure a pure population of NICM patients without the contribution of ischaemic aetiology. According to the original study design, the ischaemic aetiology of LV dysfunction was defined as, among others, 'evidence of ischaemic scar at LGE-CMR explaining the degree of LV systolic dysfunction'. By applying a more restrictive interpretation of that criterion, we excluded all patients showing a concomitant non-ischaemic and ischaemic LGE pattern. This choice allowed us to obtain more refined prognostic results at the expense of a smaller population than in the previously published analysis.

Thirdly, our registry did not include NICM patients on treatment with SGLT2i, which may potentially improve the clinical outcome of these patients, nor did it include patients with sacubitril/valsartan treatment, which potentially may reduce the risks of ventricular arrhythmia compared with angiotensin converting enzyme inhibitors/angiotensin receptor blockers therapy in HF patients.²⁵ However, during the enrolment period of our registry, the two drug concepts had not yet reached full acceptance in the cardiology community worldwide. Moreover, we did not include biomarkers of HF, such as brain natriuretic peptide, or biomarkers of myocardial fibrosis, such as serum ST2, nor did we use novel CMR techniques such as quantitative T1 mapping (due to the limited availability of T1 mapping in several study centres). In addition, an arrhythmic composite risk model, which also includes genetics, will likely be able to improve the predictive performance of our model for MAACE. Thus, caution is needed in extending the results of our registry to NICM patients with different genotypes.

Fourthly, our registry does not include genetic tests. There is a growing body of evidence in the literature that genetic mutations correlate with LGE amount and distribution.²⁶ The majority of data showed that genes can be classified into three categories according to the predominant LGE pattern distribution as subepicardial/ring-like (DMD, DSP, and FLNC), unspecific (TTN, BAG3, LMNA, and MYBPC3), and absent/rare (TNNT2, RBM20, and MYH7) with a complementary role in terms of prognostic stratification.²⁷ However, whether tissue characterization and genetic tests are faces of the same coin is still under investigation, and moreover, these tests are not still widespread in clinical practice.

Finally, even larger studies are warranted to minimize the underpowering issue related to the uneven prevalence of scarring in the different ventricular sectors. Notwithstanding these limitations, the results of our registry add important new information on how to best use LGE-CMR in risk stratification of NICM patients.

Conclusion

The incremental value of the DERIVATE Risk Score 2.0 is mainly related to the integration of LGE pattern and location in a model including

gender and continuous value of LVEF that is a unique prerogative of CMR (*Graphical Abstract*). Further randomized trials to test a CMR-guided strategy for ICD implantation vs. standard of care are now needed.

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Supplementary data

Supplementary data are available at *European Heart Journal - Cardiovascular Imaging online*.

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Data availability

The original data generated during the study are available at the following link upon request: <https://zenodo.org/records/14590431>.

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